Graphical Models and Variational Inference

Demian Wassermann, Inria Graphical Models: Discrete Inference and Learning

Introduction to DAG and their relationship with Probability Functions (Pearl)



WATSON'S CALL = TRUE

[Pearl 1987]



[Kong et al 2019]



Z: is a "Topic" W: is an observed "Word" [Blei et al 2003]

Each "box" or template represents a set of i.i.d. random variables with the same distribution

Introduction to DAG and their relationship with **Probability Functions (Pearl)**

0			
NEW	MILLION	CHILDREN	SCHOOL
FILM	TAX	WOMEN	STUDENTS
SHOW	PROGRAM	PEOPLE	SCHOOLS
MUSIC	BUDGET	CHILD	EDUCATION
MOVIE	BILLION	YEARS	TEACHERS
PLAY	FEDERAL	FAMILIES	HIGH
MUSICAL	YEAR	WORK	PUBLIC
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ACTOR	NEW	SAYS	BENNETT
\mathbf{FIRST}	STATE	FAMILY	MANIGAT
YORK	PLAN	WELFARE	NAMPHY
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ACTRESS	GOVERNMENT	CARE	ELEMENTARY
LOVE	CONGRESS	LIFE	HAITI

"Children"

"Education"

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U: is a Dirichlet or "clustering variable" Z: is a "Topic" W: is an observed "Word" [Blei et al 2003]

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	"Arts"	"Budgets"	"Children"	"Education"
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Introduction to DAG and their relationship with **Probability Functions (Pearl)** Then, we are looking for the posterior $P(U, Z | W, \alpha, \gamma) = \frac{P(U, Z, W | \alpha, \gamma)}{-1}$





 $U_{j} \sim Dirichlet(\alpha), \alpha < 1$ $Z_{i,j} \sim Multinomial(U_{j})$ $W_{i,j} \sim Multinomial\left(\gamma_{Z_{i,j}}\right)$ No Then, we are looking for the posterior $P(U, Z | W, \alpha, \gamma) = \frac{P(U, Z, W | \alpha, \gamma)}{P(U, Z | W, \alpha, \gamma)}$ α . γ) $\left(P(U, Z, W \mid \alpha, \gamma) = \Pi_j \right) P(U_j \mid \alpha) \left[\prod_i \sum_{Z_{i,j}} P(Z_{i,j} \mid U_j) P(W_{i,j} \mid Z_{i,j}, \gamma) \right] dU_j$

Introduction to DAG and their relationship with **Probability Functions (Pearl)**





Introduction to DAG and their relationship with **Probability Functions (Pearl)**





 $P(W_1, ..., W_I, Z_1, ..., Z_I, U_1, ..., U_J, \alpha, \gamma) = \prod_i \prod_j \prod_i P(W_i | Z_i, \gamma) P(Z_i | U_j) P(U_i | \alpha)$

In general, for a graphical model Graphical Model with vertices V and edges E

 $GM = (V, E), P(V) = \prod_{v \in V} P(v | Pa(v)), Pa(v) = \{v' : v' \to v \in E\}$

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WATSON'S CALL = TRUE

Here, the report and the sound are independent, given that we know if there was an earthquake: They are **conditionally** independent

P(R, S | E) = P(R | E)P(S | E) iif I(R, S, E)

GIBBON'S TESTIMONY



"Arts"	"Budgets"	"Children"	"Education"
"Arts" NEW FILM SHOW MUSIC MOVIE PLAY MUSICAL BEST ACTOR FIRST YORK OPERA	"Budgets" MILLION TAX PROGRAM BUDGET BILLION FEDERAL YEAR SPENDING NEW STATE PLAN MONEY PROCEDAMS	"Children" CHILDREN WOMEN PEOPLE CHILD YEARS FAMILIES WORK PARENTS SAYS FAMILY WELFARE MEN DEDCENT	"Education" SCHOOL STUDENTS SCHOOLS EDUCATION TEACHERS HIGH PUBLIC TEACHER BENNETT MANIGAT NAMPHY STATE DESCIDENT
OPERA THEATER ACTRESS LOVE	MONEY PROGRAMS GOVERNMENT CONGRESS	MEN PERCENT CARE LIFE	STATE PRESIDENT ELEMENTARY HAITI
2012			

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$$P(L \mid O) = \frac{\prod_{v \in V} P(v \mid Pa(v))}{\prod_{o} P(o \mid Pa(o))}, GM$$

In the case of continuous variables this is

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 $\mathbf{D}(\mathbf{I} \ \mathbf{O})$

 $= (V = L \cup O, E), \exists l \in L : o \rightarrow l \in E$

 $P(L \mid O)$





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No analytical solution, for the general case

 $P(L \mid O) = \frac{P(L, O)}{\int P(L, O) dO}$

 \mathbf{D} / \mathbf{T}

 $= (V = L \cup O, E), \exists l \in L : o \rightarrow l \in E$



$$P(L \mid O) = \frac{\prod_{v \in V} P(v \mid Pa(v))}{\prod_{o} P(o \mid Pa(o))}, GM$$

Can we approximate $P(L \mid O)$?

Relationship between a Directed Graphical Model and its Probability Law (Pearl and Paz 1985)

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 $= (V = L \cup O, E), \ \exists l \in L : o \to l \in E$

 $Q(L) \simeq P(L \mid C)$

 $\frac{P(L,O)}{\int P(L,O)dC}$

Approximations to Density Laws $\frac{Q(L)}{P(L)} \simeq P(L \mid O) = \frac{P(L, O)}{\int P(L, O) dO}$ Can we approximate $P(L \mid O)$?

- problem: how to guarantee that Q(L) is a probability law?
- Second try: cumulant approximations (changing the random L by X)

• First try: MacLaurin $Q(L) = \sum P(L = l | O) + P'(L = l | O)(l - L) + ...$



Approximations to Density Laws Can we approximate P(L|O)? $Q(L) \simeq P(L|O) = \frac{P(L,O)}{\int P(L,O)dO}$

- First try: MacLaurin $Q(L) = \sum P(L = l | O) + P'(L = l | O)(l L) + ...$ problem: how to guarantee that Q(L) is a probability law?
- Second try: cumulant approximations (changing the random L by X) $\sum \kappa_n \frac{t^n}{n!} = \kappa_1 t + \kappa_2 \frac{t^2}{2!} + \dots = \mu t + \sigma^2 \frac{t^2}{2!} + \dots$

$$\phi(t) = \log \mathbb{E}_{X \sim Q(X)}[\exp(tX)] = \sum_{X \sim Q(X)} [\exp(tX)] = \sum_{X \sim Q($$

• However, a probability law has either up to two moments, or an infinite number (Cramèr 1938)

Approximations to Density Laws $Q(L) \simeq P(L \mid O) = \frac{P(L, O)}{\int P(L, O) dO}$ Can we approximate $P(L \mid O)$?

 Other options: Edgesworth, approximations which come from this identity $\phi(t) = \log \mathbb{E}_X[\exp(itX)] = \sum \kappa_n \frac{(it)^n}{n!},$ $\psi(t) = \log \mathbb{E}_X[\exp(itX)] = \sum_{n=1}^{n} \gamma_n \frac{(it)^n}{n!}$ $\hat{\phi}(t) = \sum_{n} (\kappa_n - \gamma_n) \frac{(it)^n}{n!} + \log \psi(t)$

however, they are not guaranteed to be probability laws for finite samples.

Approximations to Density Laws $Q(L) \simeq P(L \mid O) = \frac{P(L, O)}{\int P(L, O) dO}$ Can we approximate $P(L \mid O)$?

- So? What do we do?

so we need to define the right similarity measurement D to compare distributions. And in standard Variational Inference (VI), Z is notation for O

• We choose an approximate distribution $\check{Q}_{\theta}(X)$ —replacing L by X and Oby Z for notation – from a given family, with parameters θ . Then $Q^* = Q_{\theta^*}: \theta^* = \arg\min_{\Theta} D(Q_{\theta}(X), P(X|Z))$



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aws

so we need to define the right similarity measurement D to compare distributions. And in standard Variational Inference (VI), Z is notation for O



So Which D and Q Should We Choose? $Q^* = Q_{\theta^*} : \theta^* = \arg\min_{O} D(Q_{\theta}(X), P(X|Z))$

X the latent variables and Z the observations

Let's start with "analytical" ideas: $D(Q_{\theta}(X), P(X|Z)) = \int (Q_{\theta}(x) - P(x|Z))^2 dx$

How easy is to obtain bounds and closed form solutions?

- •What does it mean for two distributions to be close in the L_2 sense?

 $\cdot Q_{\theta}(X) : X \sim \mathcal{N}(\mu, \Sigma), \theta = (\mu, \Sigma)$: This is called the Laplace approximation •Even simpler $\Sigma = \sigma^2 Id$, which boils down to $Q_{\mu}(X) = \prod_i Q_{\mu_i}(X_i)$



More Information theoretic

$$D_{KL}(Q_{\theta}(X), P(X | Z)) = \mathbb{E}_{X \sim Q_{\theta}}$$

 The Kullback-Leibler divergence is based on information theory Known formulations for common cases

•Mean field
$$Q_{\theta=\mu}(X) = \prod_i Q_{\mu_i}(X_i)$$



A Case for Mean Field KL-based VI

Journal of Artificial Intelligence Research 4 (1996) 61-76

Submitted 11/95; published 3/96

Mean Field Theory for Sigmoid Belief Networks

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Abstract

We develop a mean field theory for sigmoid belief networks based on ideas from statistical mechanics. Our mean field theory provides a tractable approximation to the true probability distribution in these networks; it also yields a lower bound on the likelihood of evidence. We demonstrate the utility of this framework on a benchmark problem in statistical pattern recognition—the classification of handwritten digits.







		1									
			1	2	3	4	5	6	7	8	9
-	0	388	2	2	0	1	3	0	0	4	0
	1	0	393	0	0	0	1	0	0	6	0
	2	1	2	376	1	3	0	4	0	13	0
->	3	$\left(\begin{array}{c} 0 \end{array} \right)$	2	4	373	0	12	0	0	6	3
	4	0	0	2	0	383	0	1	2	2	10
	5	0	2	1	13	0	377	2	0	4	1
	6	1	4	2	0	1	6	386	0	0	0
	$\overline{7}$	0	1	0	0	0	0	0	388	3	8
	8	1	9	1	7	0	7	1	1	369	4
	9	0	4	0	0	0	0	0	8	5	383



So Which D and Q Should We Choose? $Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$

X the latent variables and Z the observations

A second order information-theoretic model

• $Q_{\theta}(X): X \sim \mathcal{N}(\mu, \Sigma), \theta = (\mu, \Sigma):$ This is called the Laplace approximation



But Laplace is Better Submitted 00/00; Published 00/00

Journal of Machine Learning Research (2013)

Variational Inference in Nonconjugate Models

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- 1. Draw coefficients $\theta \sim \mathcal{N}(\mu_0, \Sigma_0)$.
- 2. For each data point n and its covariates t_n , draw its class label from



]	Yeast	Scene		
	Accuracy	Log Likelihood	Accuracy	Log Likelihoo	
akkola and Jordan (1996)	79.7%	-0.678	87.4%	-0.670	
place inference	80.1%	-0.449	89.4%	-0.259	

So Which D and Q Should We Choose? $Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$

A second order information-theoretic model

- X the latent variables and Z the observations
- $D_{KL}(Q_{\theta}(X), P(X|Z)) = \mathbb{E}_{X \sim Q_{\theta}} \left| -\log \frac{P(X|Z)}{Q_{\theta}(X)} \right| = -\int dQ_{\theta}(x)\log \frac{P(x|Z)}{Q_{\theta}(x)}$
- • $Q_{\theta}(X): X \sim \mathcal{N}(\mu, \Sigma), \theta = (\mu, \Sigma)$: This is called the Laplace approximation



So Which D Should We Choose? Finding Bounds

But our graphical model is more adapted to sample from P(X, Z) than from $P(X \mid Z).$

Then, can we find a way to efficiently minimise $D_{KL}\left(Q_{\theta}(X), \frac{P(X, Z)}{P(Z, Y)}\right)$ when, in general, we don't know the probability of "evidence" P(Z)? Let's see in the next slide....





So Which
$$D$$
 Should We C
 $D_{KL}(Q_{\theta}(X), P(X|Z)) = \mathbb{E}_{X \sim Q_{\theta}} \left[-\log And we know that$
 $\log P(Z) = \log \int dx P(x, Z) = \log \int dx P(x, Z) dx + \log dx = \log \int dx P(x, Z) dx + \log dx = \log \int dx + \log \int$

e Choose? Finding Bounds $-\log \frac{P(X|Z)}{O_{\theta}(X)} = -\int dQ_{\theta}(x)\log \frac{P(x|Z)}{Q_{\theta}(x)}$



 $\mathscr{L}(\theta)$

) before) and X our latent variables (L) $| \geq E_{X \sim Q_{\theta}} | \frac{P(X, Z)}{C} |$

 $\max \mathscr{L}(\theta)$ the Evidence Lower Bound (ELBO): $\mathscr{L}(\theta)$



So Which *D* and *Q* Should We Choose? $Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$

X the latent variables and Z the observations

A simplified second order information-theoretic model $\theta = \arg \max_{\theta} \mathscr{L}(\theta) = \mathbb{E}_{X \sim Q_{\theta}} \log$ • $Q_{\theta}(X) : X \sim \mathcal{N}(\mu, \Sigma), \theta = (\mu, \Sigma)$: This is called the Laplace approximation

$$g \frac{P(X,Z)}{Q_{\theta}(X)}$$



But Laplace is Better (they use ELBO) Submitted 00/00; Published 00/00

Journal of Machine Learning Research (2013)

Variational Inference in Nonconjugate Models

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- 1. Draw coefficients $\theta \sim \mathcal{N}(\mu_0, \Sigma_0)$.
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$$z_n \mid \theta, t_n \sim \text{Bernoulli} \left(\sigma(\theta^\top t_n)^{z_{n,1}} \sigma(-\theta^\top t_n)^{z_{n,2}} \right)$$

	Yeast		Scene		
	Accuracy	Log Likelihood	Accuracy	Log Likelihoo	
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More General Q_A $Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\theta} D(Q_{\theta}(X), P(X|Z))$

X the latent variables and Z the observations $\overline{}$

- Gaussian Processes: A measure over continuous functions where any discrete sample of the domain follows a Gaussian law. \mathcal{A}
- $P(f(x)): (f(x_1), \dots, f(x_N)) \sim$
- •Support Transformations: Q_{ℓ}

 $X \sim \mathcal{N}(\mu, \Sigma), \phi_{\theta}$ a parametric mass-preserving diffeomorphism



More General \mathcal{Q}_A $Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\Omega} D(Q_{\theta}(X), P(X|Z))$

X the latent variables and Z the observations

• Support Transformations: $Q_{\theta}(X) \triangleq N_{\mu,\Sigma}(\phi_{\theta}(X))$

 $X \sim \mathcal{N}(\mu, \Sigma), \phi_{\theta}$ a parametric mass-preserving diffeomorphism





X the latent variables and Z the observations



More General Q_A $Q^* = Q_{\theta^*} : \theta^* = \arg\min_{\varphi} D(Q_{\theta}(X), P(X|Z))$

• Support Transformations: $Q_{\theta}(X) \triangleq N_{\mu,\Sigma}(\phi_{\theta}(X)) \left| J_{\phi_{\theta}}(X) \right|$

 $\phi_{\theta}(X) \sim \mathcal{N}(\mu, \Sigma), \phi_{\theta}$ a stochastic flow or learnable diffeomorphism

[Papamakarios etal 21]



Current Problems in VI

- Scalability
- Amortization [Gershman et al 2014]
- Preservation of dependencies
- Auto-regressive models



Query 1: P(B|C) = P(C|B)P(B)/P(C)

Query 2:
$$P(A|C) = \sum_{B} P(A|B)P(B|C)$$

Amortisation, reused probability in blue



Figure 1: A Bayesian network modeling brightness constancy in visual perception, a possible invers factorization, and two of the local joint distributions that determine the inverse conditionals.

[Stuhlmüller et al 14]





Other Modern Bayesian Techniques

- Variational AutoEncoders
- Likelihood-free Inference





(b) VAE

$VAE: Z \sim N(\mu(X), \Sigma(X))$